

Riemannian geometry in elliptical distributions

Arnaud Breloy, SLSIP Workshop, Rüdesheim, October 7th 2021

Some references

“Intrinsic Cramér–Rao bounds for scatter and shape matrices estimation in CES distributions,” SPL, 2018.
Arnaud Breloy, Guillaume Ginolhac, Alexandre Renaux, Florent Bouchard

“A Riemannian Framework for Low-Rank Structured Elliptical Models,” TSP, 2021.
Florent Bouchard, Arnaud Breloy, Guillaume Ginolhac, Alexandre Renaux, Frederic Pascal

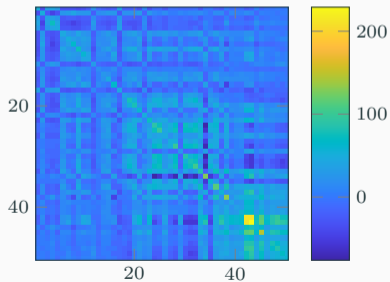
“A Tyler-Type Estimator of Location and Scatter Leveraging Riemannian Optimization,” ICASSP 2021.
Antoine Collas, Florent Bouchard, Arnaud Breloy, Guillaume Ginolhac, Chengfang Ren, Jean-Philippe Ovarlez

“Probabilistic PCA from Heteroscedastic Signals: Geometric Framework and Application to Clustering”
Antoine Collas, Florent Bouchard, Arnaud Breloy, Guillaume Ginolhac, Chengfang Ren, Jean-Philippe Ovarlez

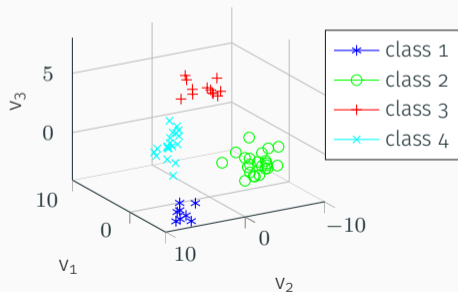
Motivations

Represent or analyze the data \mathbf{x} through some parameter θ

Example with $p \simeq 7k$ genes of $n = 63$ patients with $k = 4$ classes [Khan2001] represented by



Covariance of 50 selected genes



3 principal components

Statistical approach

“Assume $\mathbf{x} \sim f(\mathbf{x}, \boldsymbol{\theta})$, then do stuff”

- **Design** a meaningful pdf f and parameter $\boldsymbol{\theta}$
- **Analyze** model properties, performance bounds...
- **Solve** related optimization problems (MLEs, barycenters...)
- **Apply** the results to a task

Today's talk: What can Riemannian geometry bring to these steps?

Outline

X Design

- Examples of f and θ from elliptical distributions
- Remark that $\theta \in \mathcal{M} \implies$ pretext to define Riemannian tools

• Analyze

- Intrinsic Cramér-Rao bounds
- 2 examples of interesting inequalities

• Solve

- Riemannian optimization and geodesic convexity
- 2 examples where numerical stability is improved

• Apply

- Clustering with Riemannian distances

Elliptical distributions

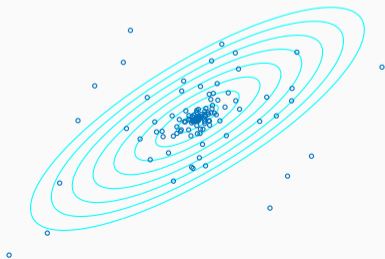
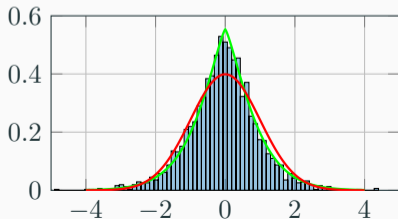
Complex elliptically symmetric distributions (CES)

$\mathbf{x} \sim \mathcal{CES}(\boldsymbol{\mu}, \boldsymbol{\Sigma}, g)$ if it has for pdf

$$f(\mathbf{x}) \propto |\boldsymbol{\Sigma}|^{-1} g((\mathbf{x} - \boldsymbol{\mu})^H \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}))$$

Scaled Gaussian distributions (CG)

$\mathbf{x}_i \sim \mathcal{CN}(\boldsymbol{\mu}, \tau_i \boldsymbol{\Sigma})$ where τ_i is unknown deterministic



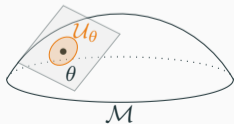
Structured parameter space as a manifold

Generally, the distribution **parameter space**, e.g.

- **Covariance matrices:** $\Sigma \in \mathcal{H}_p^{++}$
- **Product spaces:** $\{\{\tau_i\}_{i=1}^n, \mu, \Sigma\} \in (\mathbb{R}^+)^n \times \mathbb{C}^p \times \mathcal{H}_p^{++}$

turn out to be a **manifold** \mathcal{M} (locally diffeomorphic to \mathbb{R}^d , with $\dim(\mathcal{M}) = d$)

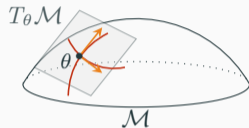
$\forall \theta \in \mathcal{M}, \exists \mathcal{U}_\theta \subset \mathcal{M}$ and $\varphi_\theta : \mathcal{U}_\theta \rightarrow \mathbb{R}^d$, diffeomorphism



Riemannian manifolds (1/2)

Tangent space $T_\theta \mathcal{M}$ at point θ

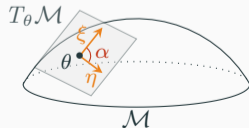
- Curve $\gamma : \mathbb{R} \rightarrow \mathcal{M}, \gamma(0) = \theta$
- Derivative: $\dot{\gamma}(0) = \lim_{t \rightarrow 0} \frac{\gamma(t) - \gamma(0)}{t}$
- Tangent space $T_\theta \mathcal{M} = \{\dot{\gamma}(0) : \gamma : \mathbb{R} \rightarrow \mathcal{M}, \gamma(0) = \theta\}$



Equip $T_\theta \mathcal{M}$ with a **Riemannian metric** $\langle \cdot, \cdot \rangle_\theta$ yields a **Riemannian manifold**

- $\langle \cdot, \cdot \rangle_\theta : (T_\theta \mathcal{M} \times T_\theta \mathcal{M}) \rightarrow \mathbb{R}$ **inner product** on $T_\theta \mathcal{M}$
(bilinear, symmetric, positive definite)
- defines length and relative positions of tangent vectors

$$\|\xi\|_\theta^2 = \langle \xi, \xi \rangle_\theta \quad \alpha(\xi, \eta) = \frac{\langle \xi, \eta \rangle_\theta}{\|\xi\|_\theta \|\eta\|_\theta}$$



Riemannian manifolds (2/2)

The Riemannian metric $\langle \cdot, \cdot \rangle_\theta$ induces **a geometry** for \mathcal{M}

Geodesics $\gamma : [0, 1] \rightarrow \mathcal{M}$

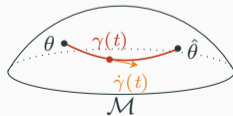
- generalizes straight lines on \mathcal{M}

- curves on \mathcal{M} with zero acceleration: $\frac{D^2\gamma}{dt^2} = 0$

defined by $(\gamma(0), \dot{\gamma}(0))$ or $(\gamma(0), \gamma(1))$

operator $\frac{D^2}{dt^2}$ depends on \mathcal{M} and $\langle \cdot, \cdot \rangle$.

Riemannian distance $\text{dist}(\theta, \hat{\theta}) = \int_0^1 \|\dot{\gamma}(t)\|_{\gamma(t)} dt$



distance = length of γ connecting θ and $\hat{\theta}$

Which metric/geometry to chose ?

The **Fisher information metric** looks like an **ideal driven by the model**

Still, we can chose **alternate metrics suited to some needs**

- Availability (**closed-form**) of theoretical objects
- Interesting **invariance** properties
- **Practical results** of the chosen task

Metric	Geodesics	Distance	Retraction	Completeness	Invariance 1	Invariance 2	Perf.
(a)	✗	✗	✓	✓	✗	✓	82%
(b)	✓	✗	✓	✓	✓	✗	86%
(c)	✓	✓	✓	✗	✗	✓	79%

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- **Analyze**

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Cramér-Rao lower bound (CRLB)

CRLB: If $\mathbf{z} \sim f(\mathbf{z}, \boldsymbol{\theta})$, then for $\hat{\boldsymbol{\theta}}$ unbiased estimator of $\boldsymbol{\theta}$ \triangleq as a vector

$$\mathbb{E} \left\{ (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})^T \right\} \succeq \mathbf{F}^{-1}(\boldsymbol{\theta}) \quad \Rightarrow \quad \text{MSE} \geq \text{Tr} \left\{ \mathbf{F}^{-1}(\boldsymbol{\theta}) \right\}$$

with the **Fisher information matrix** $\mathbf{F}(\boldsymbol{\theta}) = -\mathbb{E} \left\{ \frac{\partial^2 \ln f(\mathbf{z}, \boldsymbol{\theta})}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}^T} \middle| \boldsymbol{\theta} \right\}$

Slepian-Bangs formula: if $\mathbf{x} \sim \mathcal{CN}(\boldsymbol{\mu}(\boldsymbol{\theta}), \boldsymbol{\Gamma}(\boldsymbol{\theta}))$

$$[\mathbf{F}(\boldsymbol{\theta})]_{ij} = 2\Re \left\{ \frac{\partial \boldsymbol{\mu}^H(\boldsymbol{\theta})}{\partial \theta_i} \middle|_{\boldsymbol{\theta}} \boldsymbol{\Gamma}^{-1}(\boldsymbol{\theta}) \frac{\partial \boldsymbol{\mu}(\boldsymbol{\theta})}{\partial \theta_j} \middle|_{\boldsymbol{\theta}} \right\} + \text{Tr} \left\{ \boldsymbol{\Gamma}^{-1}(\boldsymbol{\theta}) \frac{\partial \boldsymbol{\Gamma}(\boldsymbol{\theta})}{\partial \theta_i} \middle|_{\boldsymbol{\theta}} \boldsymbol{\Gamma}^{-1}(\boldsymbol{\theta}) \frac{\partial \boldsymbol{\Gamma}(\boldsymbol{\theta})}{\partial \theta_j} \middle|_{\boldsymbol{\theta}} \right\}$$

Extension to CES in [Besson13]

“Constrained” CRLB (cCRLB)

Constraints: If elements of $\boldsymbol{\theta}$ are linked by some system

$$h_k(\theta_1, \theta_2, \dots, \theta_P) = 0, \quad k \in \llbracket 1, M \rrbracket \quad \Leftrightarrow \quad \mathbf{h}(\boldsymbol{\theta}) = \mathbf{0}$$

$\mathbf{F}(\boldsymbol{\theta})$ becomes singular \Rightarrow no proper CRLB

cCRLB: we still have [Gormango]

$$\mathbb{E} \left\{ (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})^T \right\} \succeq \mathbf{U}(\boldsymbol{\theta}) \left(\mathbf{U}^T(\boldsymbol{\theta}) \mathbf{F}(\boldsymbol{\theta}) \mathbf{U}(\boldsymbol{\theta}) \right)^{-1} \mathbf{U}^T(\boldsymbol{\theta})$$

with $\mathbf{U}(\boldsymbol{\theta})$ such that $\mathbf{H}(\boldsymbol{\theta}) \mathbf{U}(\boldsymbol{\theta}) = \mathbf{0}$ and $\mathbf{U}^T(\boldsymbol{\theta}) \mathbf{U}(\boldsymbol{\theta}) = \mathbf{I}_M$, and $\mathbf{H}(\boldsymbol{\theta}) = \left. \frac{\partial \mathbf{h}(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}^T} \right|_{\boldsymbol{\theta}}$

But what if $\theta \in \mathcal{M}$?

• Parameterization and constraints ?

- Difficult to have a system of coordinates e.g. subspaces
- Difficult (or impossible) to express constraints as $\mathbf{h}(\boldsymbol{\theta})$ e.g. PSD for \mathcal{H}_p^{++}

• Performance measure ?

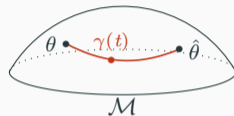
- Can we bound a Riemannian distance rather than the MSE ?
- Non-trivial function \Rightarrow no Jacobian

\rightarrow We can turn to the framework of **intrinsic CRLB** (iCRLB)

Riemannian framework of iCRLB

Definitions:

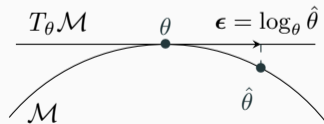
- $\theta \in \mathcal{M}$ with tangent space $T_\theta \mathcal{M}$
- $\hat{\theta} \in \mathcal{M}$ estimate of θ
- $\langle \cdot, \cdot \rangle_\theta$ chosen Riemannian metric
- $\text{dist}(\cdot, \cdot)$ induced Riemannian distance
- $\{\xi_i\}$ corresponding orthonormal basis of $T_\theta \mathcal{M}$



Error measure = $\text{dist}^2(\theta, \hat{\theta})$

Riemannian logarithm $\epsilon = \log_\theta \hat{\theta} \in T_\theta \mathcal{M}$

- Points from θ to $\hat{\theta}$ with $\|\log_\theta \hat{\theta}\|_\theta^2 = \text{dist}^2(\theta, \hat{\theta})$
- Would be " $\hat{\theta} - \theta$ " in the Euclidean setup
- In coordinates $[\epsilon]_i = \langle \log_\theta \hat{\theta}, \xi_i \rangle_\theta$



Error vector = $\log_\theta \hat{\theta}$

Fisher information metric/matrix

Fisher information metric For $f(\{\mathbf{x}_k\}; \theta)$ p.d.f. parameterized by $\theta \in \mathcal{M}$

$$\langle \xi, \xi \rangle_{\theta}^{\text{FIM}} = -\mathbb{E} \left[\left. \frac{d^2}{dt^2} \ln f(\{\mathbf{x}_k\}; \theta + t\xi) \right|_{t=0} \right]$$

Fisher information matrix represented in coordinates $\{\xi_i\}$ by

$$[\mathbf{F}]_{ij} = \langle \xi_i, \xi_j \rangle_{\theta}^{\text{FIM}}$$

Remarks

- $\langle \cdot, \cdot \rangle_{\theta}^{\text{FIM}}$ defines a metric for $T_{\theta}\mathcal{M} \Rightarrow$ **information geometry** for \mathcal{M}
- Error measured from $\langle \cdot, \cdot \rangle_{\theta}$, which can be different

Intrinsic CRLB

Intrinsic CRLB (iCRLB)

[Smith05, Boumal14]

Assuming model $f(\{\mathbf{x}_k\}; \boldsymbol{\theta})$ and unbiased estimator $\hat{\boldsymbol{\theta}}$, we have

$$\mathbb{E} \left[(\log_{\theta} \hat{\boldsymbol{\theta}}) (\log_{\theta} \hat{\boldsymbol{\theta}})^H \right] \succeq \mathbf{F}^{-1} - \underbrace{\frac{1}{3} (\mathbf{F}^{-1} \mathbf{R}_m (\mathbf{F}^{-1}) + \mathbf{R}_m (\mathbf{F}^{-1}) \mathbf{F}^{-1})}_{\text{Riemannian curvature terms (cf. [Boumal14, Eq.6.6])}} + \mathcal{O}(\lambda_{\max}(\mathbf{F}^{-1})^{2+1/2})$$

Remarks

- \mathbf{F}^{-1} depends on $\langle \cdot, \cdot \rangle_{\theta} \Rightarrow$ iCRLB indeed changes w.r.t. d (“ $(\cdot)^{-1}$ ” inverse of a tensor (defined w.r.t. a metric))
- Bias terms + more about curvature in [Smith05]
- Neglecting the curvature terms, we have in trace $\mathbb{E} \left\{ \text{dist}^2(\hat{\boldsymbol{\theta}}, \boldsymbol{\theta}) \right\} \geq \text{Tr} \left\{ \mathbf{F}^{-1} \right\}$

Wrapping up

iCRLB cooking recipe

1. Compute $\langle \xi, \xi \rangle_{\theta}^{\text{FIM}} = -\mathbb{E} \left[\frac{d^2}{dt^2} \ln f(\{\mathbf{x}_k\}; \theta + t\xi) \Big|_{t=0} \right]$ and polarization for $\langle \xi_i, \xi_j \rangle_{\theta}^{\text{FIM}}$
2. Chose the error metric $\langle \cdot, \cdot \rangle_{\theta} \rightarrow \begin{cases} \text{error distance dist} \\ \text{orthonormal basis } \{\xi_i\} \text{ of } T_{\theta}\mathcal{M} \end{cases}$
3. Compute the Fisher information matrix: $[\mathbf{F}]_{ij} = \langle \xi_i, \xi_j \rangle_{\theta}^{\text{FIM}}$
4. Bound the expected distance as $\mathbb{E} \left\{ \text{dist}^2(\hat{\theta}, \theta) \right\} \geq \text{Tr} \{ \mathbf{F}^{-1} \}$

Interest?

- Bounding other distances: neat formulas, reveals unexpected things (intrinsic bias)
- Parameterization from $T_{\theta}\mathcal{M} \rightarrow$ useful even in the Euclidean case!

Example 1: iCRLB for covariance matrix estimation in CES (1/2)

Model $\mathbf{x} \sim \mathcal{CES}(\mathbf{0}, \Sigma, g)$ with pdf $f(\mathbf{x}) \propto |\Sigma|^{-1} g(\mathbf{x}^H \Sigma^{-1} \mathbf{x})$, and representation

$$\mathbf{x} \stackrel{d}{=} \sqrt{Q} \Sigma^{1/2} \mathbf{u} \quad \text{with} \quad \begin{cases} \mathbf{u} \text{ uniformly distributed on the unit sphere } \mathbf{u} \sim \mathcal{U}(\mathbb{CS}^p) \\ Q \text{ independent modular variate, pdf related to } g \end{cases}$$

Manifold $\Sigma \in \mathcal{H}_p^{++}$ with tangent space $T_{\Sigma} \mathcal{H}_p^{++} = \mathcal{H}_p$
 (Hermitian pd matrices) (Hermitian matrices)

Error metric: “natural” Riemannian metric and distance for \mathcal{H}_p^{++}

$$\langle \xi_i, \xi_j \rangle_{\Sigma} = \text{Tr} \{ \Sigma^{-1} \xi_i \Sigma^{-1} \xi_j \} \quad \text{inducing} \quad \text{dist}_{\mathcal{H}_p^{++}}^2(\Sigma, \hat{\Sigma}) = \|\log \Sigma^{-1/2} \hat{\Sigma} \Sigma^{-1/2}\|_F^2$$

Example 1: iCRLB for covariance matrix estimation in CES (2/2)

Fisher information metric for CES

aka “affine invariant”

Let $\{\mathbf{x}_i\}_{i=1}^n$ in \mathbb{C}^p with $\mathbf{x} \sim \mathcal{CES}(\mathbf{0}, \Sigma, g)$, then

$$\langle \xi_i, \xi_j \rangle_{\Sigma}^{\text{FIM}} = n\alpha_g \text{Tr} \left\{ \Sigma^{-1} \xi_i \Sigma^{-1} \xi_j \right\} + n\beta_g \text{Tr} \left\{ \Sigma^{-1} \xi_i \right\} \text{Tr} \left\{ \Sigma^{-1} \xi_j \right\}$$

with $\alpha_g = 1 - \frac{\mathbb{E}[\mathcal{Q}^2 \phi'(\mathcal{Q})]}{M(M+1)}$ and $\beta_g = \alpha - 1$ using $\phi(t) = g'(t)/g(t)$

iCRLB for Σ

Let $\{\mathbf{z}_i\}_{i=1}^n$ in \mathbb{C}^p with $\mathbf{z} \sim \mathcal{CES}(\mathbf{0}, \Sigma, g)$

$$\mathbb{E} \left[\text{dist}_{\mathcal{H}_p^{++}}^2 \left(\hat{\Sigma}, \Sigma \right) \right] \geq \frac{1}{n} \left(\frac{p^2 - 1}{\alpha_g} + \frac{1}{\alpha_g(p+1) - p} \right)$$

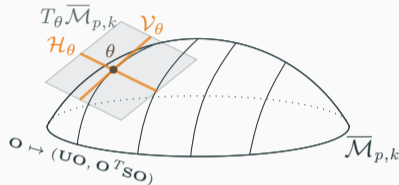
Example 2: spiked model (PCA) in CES

Model: $\mathbf{x} \sim \mathcal{CES}(\mathbf{0}, \mathbf{H} + \mathbf{I}, g)$, with $\mathbf{H} \in \mathcal{H}_{p,k}^+$
(H-psd of rank k)

Manifold: $\mathbf{H} = \mathbf{U}\Sigma\mathbf{U}^H \in \mathcal{H}_{p,k}^+$ as $(\text{St}(p, k) \times \mathcal{H}_k^{++})/\mathcal{U}_k$

Error metric:

$$\langle \bar{\xi}, \bar{\eta} \rangle_{\bar{\theta}} = \underbrace{\Re(\text{Tr}(\xi_{\mathbf{U}}^H (\mathbf{I}_p - \frac{1}{2}\mathbf{U}\mathbf{U}^H)\eta_{\mathbf{U}}))}_{\text{canonical on } \text{St}(p, k)} + \underbrace{\alpha \text{Tr}(\Sigma^{-1}\xi_{\Sigma}\Sigma^{-1}\eta_{\Sigma}) + \beta \text{Tr}(\Sigma^{-1}\xi_{\Sigma})\text{Tr}(\Sigma^{-1}\eta_{\Sigma})}_{\text{affine invariant on } \mathcal{H}_k^{++}}$$



iCRLB for subspace

Let $\{\mathbf{z}_i\}_{i=1}^n$ in \mathbb{C}^p with $\mathbf{z} \sim \mathcal{CES}(\mathbf{0}, \mathbf{U}\text{diag}(\{\sigma_r\}_{r=1}^k)\mathbf{U}^H + \mathbf{I}, g)$

$$\mathbb{E} \left[\text{dist}_{\mathcal{G}_{p,k}}^2 \left(\text{span}(\hat{\mathbf{U}}), \text{span}(\mathbf{U}) \right) \right] \geq \frac{p-k}{n\alpha_g} \sum_{r=1}^k \frac{1+\sigma_r}{\sigma_r^2}$$

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- Remark that $\theta \in \mathcal{M} \implies$ pretext to define Riemannian tools



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- 2 examples of interesting inequalities



- ~~**Solve**~~

- Riemannian optimization and geodesic convexity
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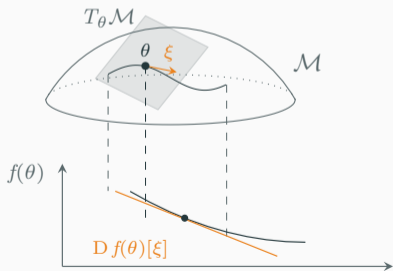
- **Apply**

- Clustering with Riemannian distances

Riemannian optimization

$$\underset{\theta \in \mathcal{M}}{\text{minimize}} \quad f(\theta)$$

Riemannian optimization: a framework for optimization on \mathcal{M} equipped with $\langle \cdot, \cdot \rangle$.



Descent direction of f at θ :

$$\xi \in T_{\theta}\mathcal{M}, \quad Df(\theta)[\xi] < 0$$

Riemannian gradient of f at θ :

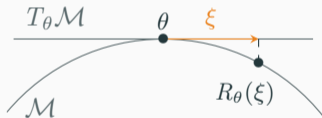
$$\langle \text{grad } f(\theta), \xi \rangle_{\theta} = Df(\theta)[\xi]$$

Riemannian optimization

[?]

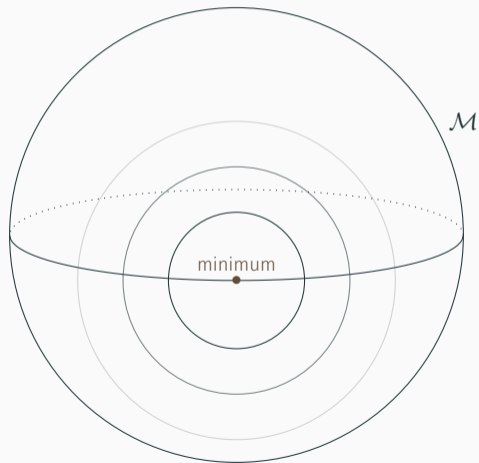
Main ingredients

- Descent direction: $\xi \in T_{\theta}\mathcal{M}$ so that $\langle \text{grad } f(\theta), \xi \rangle_{\theta} < 0$
- **Retraction** of ξ on \mathcal{M} (smooth mapping)

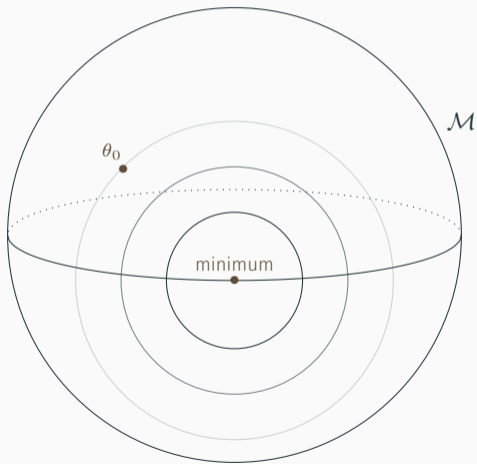


Flexibility: metric, retraction, descent method (gradient, conjugate gradient, BFGS...)

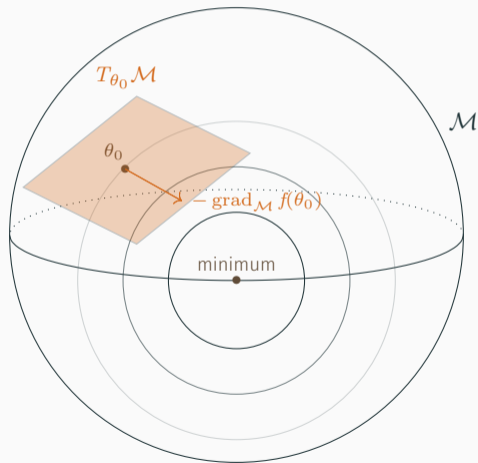
Example: Riemannian gradient descent $\theta_{i+1} = R_{\theta_i}(-t_i \text{grad } f(\theta_i))$



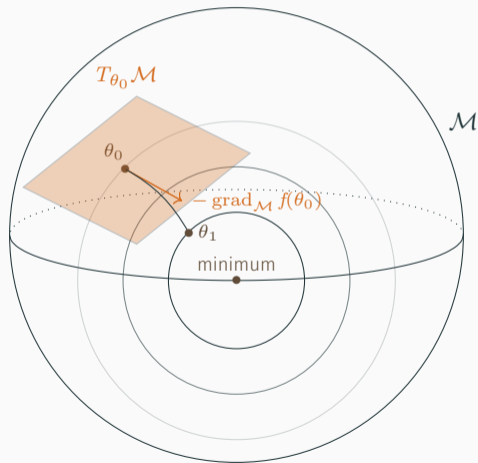
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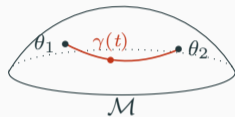


Geodesic convexity (g -convexity)

f is **g -convex** if $\forall \theta_1, \theta_2 \in \mathcal{M}$, f is convex on geodesic $\gamma(t)$, i.e

$$f(\gamma(t)) \leq t f(\theta_1) + (1 - t) f(\theta_2)$$

If so, then any local minimizer is a global minimizer.



Example: CES log-likelihoods

$$\mathcal{L}(\Sigma) = \sum_{i=1}^n \ln g(\mathbf{x}_i^H \Sigma^{-1} \mathbf{x}_i) + n \ln |\Sigma|$$

are g -convex following the geodesics $\Sigma(t) = \Sigma_1^{1/2} \left(\Sigma_1^{-1/2} \Sigma_2 \Sigma_1^{-1/2} \right)^t \Sigma_1^{1/2}$

has been useful to prove uniqueness of (regularized) M -estimators

Example 1: robust mean and covariance estimation (1/3)

Jointly estimate $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ for $\mathbf{x} \sim \mathcal{CES}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

M-estimators of location and scatter

$$\boldsymbol{\mu} = \left(\sum_{i=1}^n u_1(t_i) \right)^{-1} \sum_{i=1}^n u_1(t_i) \mathbf{x}_i \quad \boldsymbol{\Sigma} = \frac{1}{n} \sum_{i=1}^n u_2(t_i) (\mathbf{x}_i - \boldsymbol{\mu})(\mathbf{x}_i - \boldsymbol{\mu})^H$$

where $t_i \triangleq (\mathbf{x}_i - \boldsymbol{\mu})^H \boldsymbol{\Sigma}^{-1} (\mathbf{x}_i - \boldsymbol{\mu})$, and u_1, u_2 respect conditions in [Maronna76]

Tyler's estimator

$$\boldsymbol{\mu} = \left(\sum_{i=1}^n \frac{1}{\sqrt{t_i}} \right)^{-1} \sum_{i=1}^n \frac{\mathbf{x}_i}{\sqrt{t_i}} \quad \boldsymbol{\Sigma} = \frac{1}{n} \sum_{i=1}^n \frac{(\mathbf{x}_i - \boldsymbol{\mu})(-\boldsymbol{\mu})^H}{t_i}$$

Example 1: robust mean and covariance estimation (2/3)

Alternatively when $\boldsymbol{\mu} = \mathbf{0}$: Tyler's estimator \Leftrightarrow **MLE for scaled Gaussian** $\mathbf{x}_i \sim \mathcal{CN}(\mathbf{0}, \tau_i \boldsymbol{\Sigma})$

Transposed to non-zero mean $\mathbf{x}_i \sim \mathcal{CN}(\boldsymbol{\mu}, \tau_i \boldsymbol{\Sigma})$

$$\underset{\boldsymbol{\mu}, \{\tau_i\}_{i=1}^n, \boldsymbol{\Sigma}}{\text{maximize}} \sum_{i=1}^n \left[\ln |\tau_i \boldsymbol{\Sigma}| + \frac{(\mathbf{x}_i - \boldsymbol{\mu})^H \boldsymbol{\Sigma}^{-1} (\mathbf{x}_i - \boldsymbol{\mu})}{\tau_i} \right]$$

yields

$$\boldsymbol{\mu} = \left(\sum_{i=1}^n \frac{1}{t_i} \right)^{-1} \sum_{i=1}^n \frac{\mathbf{x}_i}{t_i} \qquad \boldsymbol{\Sigma} = \frac{1}{n} \sum_{i=1}^n \frac{(\mathbf{x}_i - \boldsymbol{\mu})(-\boldsymbol{\mu})^H}{t_i}$$

slightly different but fixed-point iterations diverge in practice!

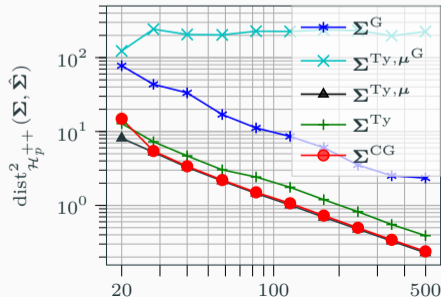
Example 1: robust mean and covariance estimation (3/3)

Product manifold $\mathcal{M}_{p,n} \in \mathbb{C}^p \times (\mathbb{R}_*^+)^n \times \mathcal{SH}_p^{++}$ with **decoupled metric**
 $(\mathcal{H}_p^{++} \cap \det = 1)$

$$\langle \xi, \eta \rangle_{\theta}^{\mathcal{M}_{p,n}} = \underbrace{\Re\{\xi_{\mu}^H \eta_{\mu}\}}_{\text{canonical on } \mathbb{C}^p} + \underbrace{(\tau^{\odot -1} \odot \xi_{\tau})^T (\tau^{\odot -1} \odot \eta_{\tau})}_{\text{canonical on } (\mathbb{R}_*^+)^n} + \underbrace{\text{Tr}(\Sigma^{-1} \xi_{\Sigma} \Sigma^{-1} \eta_{\Sigma})}_{\text{Natural Riem. on } \mathcal{SH}_p^{++}}$$

And resulting:

- Riemannian gradient descent
- Surprisingly stable and accurate estimator
- Still... slow convergence
- Faster with information geometry to appear!



Example 2: robust estimator for spiked models in CES (1/2)

Spiked Tyler's estimator

$$\begin{aligned} & \underset{\Sigma}{\text{minimize}} && \frac{p}{n} \sum_{i=1}^n \ln(\mathbf{x}_i^H \Sigma^{-1} \mathbf{x}_i) + \ln |\Sigma| \\ & \text{subject to} && \Sigma = \mathbf{H} + \sigma^2 \mathbf{I}, \text{ with } \mathbf{H} \in \mathcal{H}_{p,k}^+ \end{aligned}$$

Existing **MM algorithm** [Sun16]

1. Usual fixed point iteration

$$\Sigma_{t+1/2} = \frac{p}{n} \sum_{i=1}^n \frac{\mathbf{x}_i \mathbf{x}_i^H}{\mathbf{x}_i^H \Sigma_t^{-1} \mathbf{x}_i}$$

2. Projection on the structured set

$$\Sigma_{t+1} = \mathcal{P}_{\mathcal{H}_{p,k}^+}(\Sigma_{t+1/2})$$

where $\mathcal{P}_{\mathcal{H}_{p,k}^+}$ averages the last $p - k$ eigenvalues (SVD)

can diverge with small n

Example 2: robust estimator for spiked models in CES (2/2)

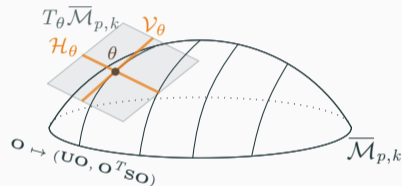
Riemannian optimization for

$$\underset{\mathbf{H} \in \mathcal{H}_{p,k}^+}{\text{minimize}} \quad \mathcal{L}_{\text{Ty}}(\mathbf{H} + \mathbf{I})$$

with $\mathbf{H} = \mathbf{U}\Sigma\mathbf{U}^H \in (\text{St}(p, k) \times \mathcal{H}_k^{++})/\mathcal{U}_k$

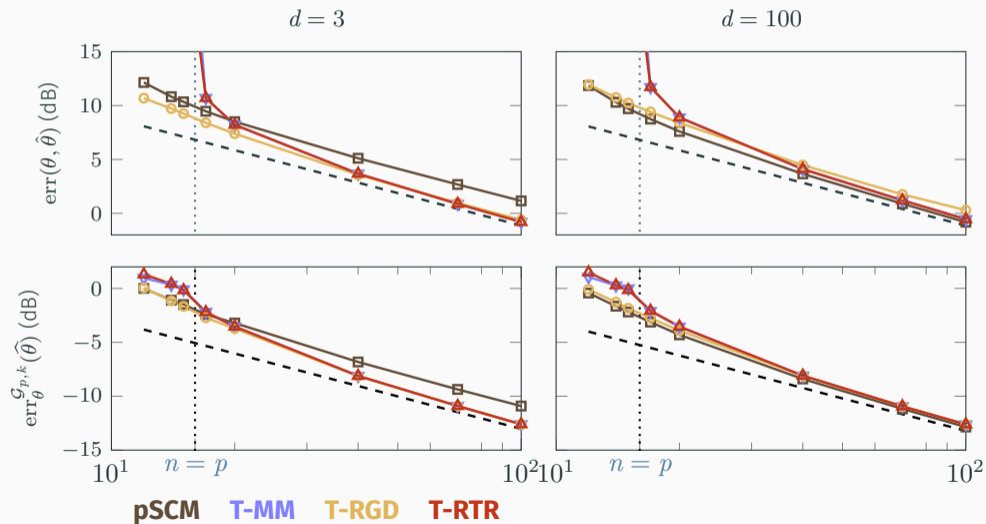
using the metric

$$\langle \bar{\xi}, \bar{\eta} \rangle_{\bar{\theta}} = \underbrace{\Re(\text{Tr}(\xi_{\mathbf{U}}^H (\mathbf{I}_p - \frac{1}{2} \mathbf{U}\mathbf{U}^H) \eta_{\mathbf{U}}))}_{\text{canonical on } \text{St}(p, k)} + \underbrace{\alpha \text{Tr}(\Sigma^{-1} \xi_{\Sigma} \Sigma^{-1} \eta_{\Sigma}) + \beta \text{Tr}(\Sigma^{-1} \xi_{\Sigma}) \text{Tr}(\Sigma^{-1} \eta_{\Sigma})}_{\text{affine invariant on } \mathcal{H}_k^{++}}$$



→ Riemannian **gradient descent (T-RGD)** and **trust region (T-RTR)** algorithms

Numerical illustrations: t -distribution $p = 16, k = 8, \text{SNR} \simeq 15\text{dB}$



Outline

• Design

- Examples of f and θ from elliptical distributions
- Remark that $\theta \in \mathcal{M} \implies$ pretext to define Riemannian tools

• Analyze

- Intrinsic Cramér-Rao bounds
- 2 examples of interesting inequalities

• Solve

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- 2 examples where numerical stability is improved

✗ Apply

- Clustering with Riemannian distances

Clustering problem

Mixture model: observations follow $(\mathbf{x}|\text{class } k) \sim f(x, \theta_k)$ with K possible classes

Clustering: from **unlabeled** data $\{\mathbf{x}_i\}_{i=1}^n$ find the partition $\{\{\mathbf{x}_i^k\}_{i=1}^{n_k}\}_{k=1}^K$

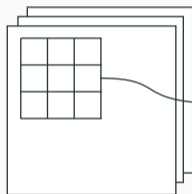
Issues:

- Statistical ideal would be the **EM algorithm** → no time for that!
- More accurate model could involve f_k → need for robustness to mismatches
- Elements in θ_k might be non-discriminating

A standard solution is to cluster intermediate **features**
(aka descriptors)

Feature clustering pipeline with a geometric twist

Step 1: **sliding window**



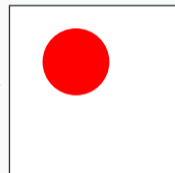
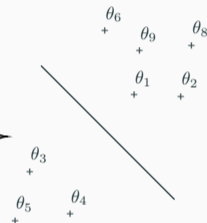
$\{x_i\}_{i=1}^N$

minimize $\mathcal{L}(\theta, x_1, \dots, x_n)$
 $\theta \in \mathcal{M}$

θ characterizes one pixel!

Step 2: **parameter estimation**

Step 3: **parameter clustering**



Riemannian approach for $\theta \in \mathcal{M}$: **transpose** clusterings algorithm using

- **Information geometry** of model $\mathbf{x} \sim f(x, \theta)$
- **Distances** $\text{dist}^2(\theta_i, \theta_j)$ and **Riemannian means** $\underset{\theta}{\text{argmin}} \sum_{i=1}^j \text{dist}^2(\theta, \theta_i)$

computed with Riem. opt.

An example on Indian pines data set

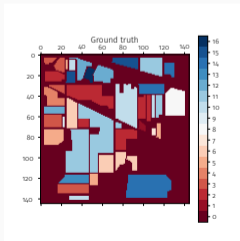
Plain K -means++, compared to Riemmanian counterparts from two models:

Centered Gaussian (GMM)

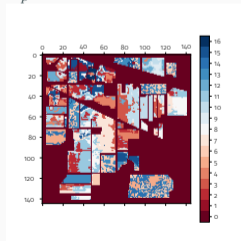
- $\mathbf{x}_i \sim \mathcal{CN}(\mathbf{0}, \Sigma)$
- $\theta =$ covariance matrix Σ
- Natural distance $\text{dist}_{\mathcal{H}_p^{++}}$

Probabilistic PCA with SG signals

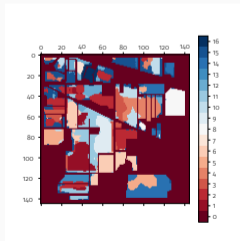
- $\mathbf{x}_i \sim \mathcal{CN}(\mathbf{0}, \tau_i \mathbf{U}\mathbf{U}^H + \mathbf{I})$
- $\theta =$ subspace $\text{span}(\mathbf{U}) +$ textures $\{\tau_i\}_{i=1}^n$
- Decoupled distance on $\text{Gr}_p^k \times \mathbb{R}^n$



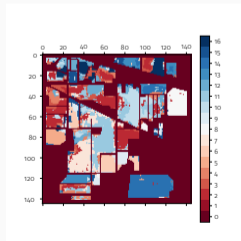
Ground truth



\mathbb{R}^p : OA = 31.2%



\mathcal{H}_p^{++} : OA = 45.2%



$\text{Gr}_p^k \times \mathbb{R}^n$: OA = 47.2%

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Perspectives in regularization

- **Intrinsic bias** $\mathbf{b}(\hat{\theta}) = \mathbb{E} \left[\log_{\theta} \hat{\theta} \right] \rightarrow$ **counter-intuitive bias-variance paradigm**

$$\mathbf{b}(\Sigma_{\text{SCM}}) = \mathbb{E} [\log_{\Sigma} \Sigma_{\text{SCM}}] = \Sigma^{1/2} \mathbb{E} \left[\log(\Sigma^{-1/2} \Sigma_{\text{SCM}} \Sigma^{-1/2}) \right] \Sigma^{1/2} = \text{not zero!}$$

- **Geodesic shrinkage** or not?

$$\Sigma_g(t) = \mathbf{T}_1^{1/2} \left(\mathbf{T}_1^{-1/2} \hat{\Sigma} \mathbf{T}_1^{-1/2} \right)^t \mathbf{T}_1^{1/2} \quad \text{versus} \quad \Sigma_L(t) = t \hat{\Sigma} + (1-t) \mathbf{T}$$

- Can we do **Stein's type** regularization (shrink eigenvalues) for $\mathbb{E} \left[\text{dist}_{\mathcal{H}_p^{++}}(\tilde{\Sigma}, \Sigma) \right]$?

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“Probabilistic PCA from Heteroscedastic Signals: Geometric Framework and Application to Clustering”

Antoine Collas, Florent Bouchard, Arnaud Breloy, Guillaume Ginolhac, Chengfang Ren, Jean-Philippe Ovarlez